

From rough to multifractal multidimensional volatility: a multidimensional Log S-fBM model

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Overview

1. Introduction and context
2. The Log S-fBM model
3. The multidimensional Log S-fBM model
4. Conclusion

Introduction and context

Introduction: The Log S-fBM model unifies rough and multifractal volatility

Multifractal Volatility Models

- Scaling property:

$$\mathbb{E}(|r(t, \Delta t)|^q) \sim \Delta t^{\zeta_r(q)}$$

- Log-correlated log volatility increments:

$$\sim -\lambda^2 \log(\tau/T)$$

- Non-linear $\zeta_r(q)$

Rough Volatility Models

- Fractional Gaussian models for the log volatility process (fBM, fOU, ...)

- Log volatility increments with correlation:

$$\sim \tau^{2H}, H \in \left] 0, \frac{1}{2} \right[$$

The Log-SfBM Model

The Log S-fBM model

The S-fBM process

Proposition [Wu *et al.*(2022)]

A *S-fBM process* $\omega_{H,T}$ is a stationary gaussian process with:

$$\forall (t, \tau) \in \mathbb{R}_+, \quad \text{Cov}(\omega_{H,T}(t + \tau), \omega_{H,T}(t)) = \frac{\lambda^2}{2H(1 - 2H)} \left(1 - \left(\frac{\tau}{T}\right)^{2H}\right) \mathbb{1}_{\{|\tau| \leq T\}}. \quad (1)$$

where $H \in]0, \frac{1}{2}[$.

Key properties:

- Fractional regularity controlled by H .
- λ^2 is the intermittency parameter, identifiable when $H \sim 0$ as $\lambda^2 \sim 2H \text{Var}(\omega_{H,T}(t))$

The Log S-fBM model: Unifying rough and multifractal volatility

The Log S-fBM random measure $M_{H,T}$ is defined as: $M_{H,T}(dt) = \exp(\omega_{H,T}(t))dt$.

Proposition [Wu et al.(2022)]

There exists a multifractal random measure $\tilde{M}_T(dt)$ such that:

$$M_{H,T}(dt) \xrightarrow[H \rightarrow 0]{w} \tilde{M}_T(dt) \quad (2)$$

This allows to derive a stochastic volatility model, the **Log S-fBM model** valid in the rough and multifractal regimes:

Definition [Wu et al.(2022)]

The **Log S-fBM model** is of the form $dX_t = \sqrt{\frac{M_{H,T}(dt)}{dt}} dB_t$, where $M_{H,T} \perp\!\!\!\perp B$.

The small intermittency approximation

By definition: $\forall I \subset \mathbb{R}_+$, $M_{H,T}(I) := \int_I e^{\omega_{H,T}(t)} dt$ and considering $\Omega_{H,T}(I) = \frac{1}{\lambda} \int_I (\omega_{H,T}(u) - \mathbb{E}[\omega_{H,T}(u)]) du$.

Proposition [Wu et al.(2022)]

Let t_1, \dots, t_n be n arbitrary times and $\Delta > 0$. The following holds:

$$\mathbb{E} \left[\ln \left(\frac{M_{H,T,\Delta}(t_1)}{\Delta} \right) \cdots \ln \left(\frac{M_{H,T,\Delta}(t_n)}{\Delta} \right) \right] = \lambda^n \Delta^{-n} \mathbb{E} [\Omega_{H,T,\Delta}(t_1) \cdots \Omega_{H,T,\Delta}(t_n)] + o(\lambda^n), \quad (3)$$

where $\Omega_{H,T,\Delta}(t) = \Omega_{H,T}([t, t + \Delta])$ and $M_{H,T,\Delta}(t) = M_{H,T}([t, t + \Delta])$.

Given this result, the calibration is performed through the so called **Generalized method of moments (GMM)**.

A particular multivariate Log S-fBM model: The Nested stationary fractional factor model (N-SfFM) (Zarhali *et al.*)

- Zarhali *et al.* propose as a multivariate stochastic volatility model using the Log S-fBM process:

$$\forall i \in \{1, \dots, N\}, \quad \begin{cases} dx_t^i = \beta_i df_t + de_t^i \\ df_t = e^{\frac{\Omega_t}{2}} dB_t \\ de_t^i = \sigma_i e^{\frac{\gamma_i \Omega_t + \omega_t^i}{2}} dB_t^i \end{cases}$$

- We build upon the intuitions behind the Nested factor model of Bouchaud & Chicheportiche(2015), reproducing:
 - A non elliptical joint distribution of asset returns
 - Non linear dependences between assets

The multidimensional Log S-fBM model

The multidimensional S-fBM process (mS-fBM)

From now on, we consider the two symmetric matrices:

- $\mathbf{H} := (H_{i,j})_{1 \leq i,j \leq d} \in]0, \frac{1}{2}[^{d \times d}$ the coHurst matrix
- $\boldsymbol{\xi} := (\xi_{i,j})_{1 \leq i,j \leq d} \in \mathbb{R}^{d \times d}$ namely the cointermittency matrix with diagonals $\xi_{i,i} := \lambda_i^2$

such that:

$$\mathcal{H}_1 : \quad \boldsymbol{\xi} \text{ is positive definite} \quad (4)$$

$$\mathcal{H}_2 : \quad \mathbf{H} \text{ is conditionally negative definite, i.e: } \forall c \in \mathbb{R}^d, \sum_{i,j=1}^d c_i c_j H_{ij} \leq 0, \text{ such that } \sum_{i=1}^d c_i = 0 \quad (5)$$

are satisfied.

Proposition [Zarhali *et al.*(2026)]

A *mS-fBM process* $((\omega_i(t))_t, i \in \llbracket 1, d \rrbracket)$ is stationary and gaussian with:

$$\forall (i, j) \in \llbracket 1, d \rrbracket^2, \quad \text{Cov}(\omega_i(t), \omega_j(s)) \sim \xi_{i,j} \left[C_{i,j} - \left(\frac{\tau}{T} \right)^{2H_{i,j}} D_{i,j} \right] \mathbb{1}_{\{\tau \leq T\}}. \quad (6)$$

where $\tau = |t - s|$, $C_{i,j}$ and $D_{i,j}$ depend on $H_{i,j}$ and $\bar{H}_{i,j} := \frac{H_{i,i} + H_{j,j}}{2}$.

\Rightarrow As in 1D, $H_{i,j}$ represents the decorrelation speed parameter.

mLog S-fBM model: Unifying rough and multifractal multidimensional volatility

Let \mathbf{M}_H , an \mathbb{R}^d valued random measure defined as:

$$(\mathbf{M}_H(dt))_j = \begin{cases} M_j(dt), & \text{if } H_{j,j} \neq 0 \\ \tilde{M}^j(dt), & \text{if } H_{j,j} = 0 \end{cases} \quad (7)$$

$M_j(dt) := \exp(\omega_j(t)) dt$ and \tilde{M}^j is the i^{th} marginal of a \mathbb{R}^d valued multifractal random measure $\tilde{\mathbf{M}}$.

Theorem [Weak continuity and limit $\|\mathbf{H}\| \rightarrow 0$, Zarhali *et al.*(2026)]

- The sequence of measures $(\mathbf{M}_{H_n})_n$ converges weakly to \mathbf{M}_H as $\|\mathbf{H}_n - \mathbf{H}\| \rightarrow 0$.
- The measure \mathbf{M}_H converges weakly to $\tilde{\mathbf{M}}$ as $\|\mathbf{H}\| \rightarrow 0$.

\Rightarrow This unifies multidimensional rough and multifractal volatility.

mLog S-fBM process and the small intermittency approximation

Definition [Zarhali *et al.*(2026)]

A *mLog S-fBM* is the \mathbb{R}^d valued process $(\mathbb{X}_t)_t := ((X_t^i)_t, i \in \llbracket 1, d \rrbracket)$ with:

$$\begin{cases} dX_t^i = \sqrt{\frac{(\mathbf{M}_H(dt))_i}{dt}} dB_t^i, \\ X_0^i = x_0^i. \end{cases} \quad (8)$$

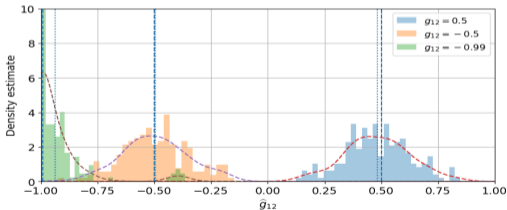
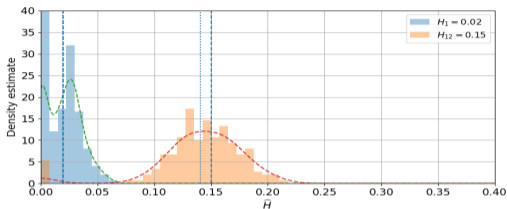
where $((B_t^i)_t, i \in \llbracket 1, d \rrbracket) \perp\!\!\!\perp \mathbf{M}_H$.

Corollary [Small intermittency approximation, case $d=2$]

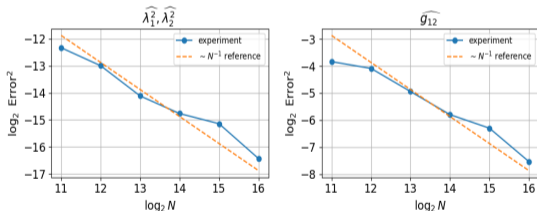
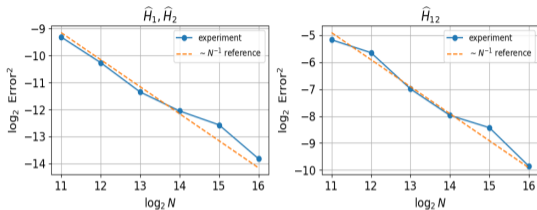
$$\forall (I, J) \subset \mathbb{R}_+^2, \quad \mathbb{E} \left(\ln \left(\frac{M_i(I)}{|I|} \right) \ln \left(\frac{M_j(J)}{|J|} \right) \right) = \lambda_i \lambda_j \mathbb{E} \left(\frac{\Omega_i(I)}{|I|} \frac{\Omega_j(J)}{|J|} \right) + o \left(\|\Lambda_{i,j}\|^2 \right) \quad (9)$$

where $(i, j) \in \llbracket 1, d \rrbracket$, $\Lambda_{i,j} = \begin{pmatrix} \lambda_i \\ \lambda_j \end{pmatrix}$ and $\Omega_i(I) = \frac{1}{\lambda_i} \int_I (\omega_i(u) - \mathbb{E}[\omega_i(u)]) du$

Estimation Robustness

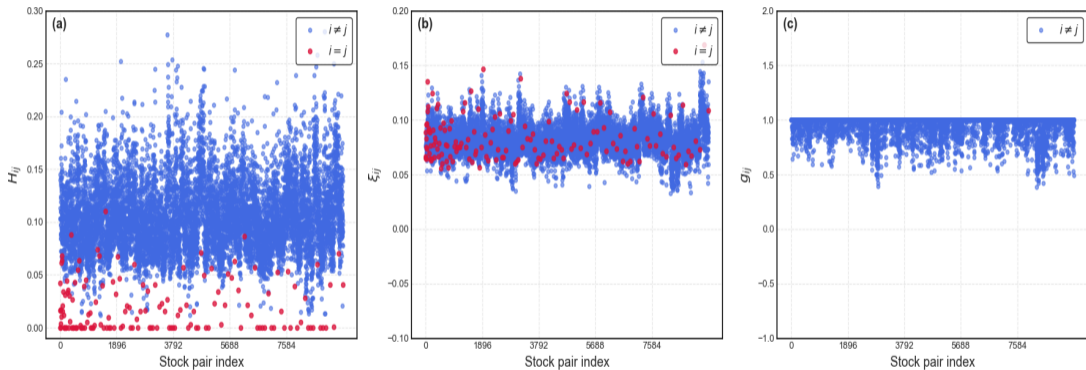


Empirical distributions of $H_{i,j}$ (**top**) and $g_{i,j}$ (**bottom**) from 200 synthetic samples.



(**Blue**) Log-standard deviation of the calibrated parameters; (**orange**) corresponding linear fits.

Calibration on S&P500 data



Scatter plots for all the selected assets of the S&P500 of: the Hursts and coHurst exponents **(a)**, the intermittency and cointermittency coefficients **(b)** and the cointermittency correlation coefficients **(c)**.

Conclusion and research directions

The purpose of this talk was to demonstrate the construction of multivariate dynamics using the Log S-fBM model:

- The **mLog S-fBM** is richer and tractable with a systematic calibration method
- Many directions remain: understand the structure of the co-Hurst matrix, investigate the correlation scale invariance ...

Thanks for you attention



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